

Mascot-Num Presentation

PhD student : Mathis Deronzier

Supervised by : François Bachoc, Olivier Roustant, and Andrés F. López-Lopera

Institut de Mathématiques de Toulouse, Université Toulouse

Regression under Monotonicity and Fairness Constraints

- (1) Block-Additive Gaussian Processes under Monotonicity Constraints
Statistics and Computing - M. D., Andrés F. López-Lopera, François Bachoc, Olivier Roustant, and Jérémy Rohmer
- (2) On the Nonconvexity of Push-Forward Constraints and Its Consequences in Machine Learning
SIMODS - Lucas De Lara, M. D., Alberto González-Sanz, and Virgil Foy
- (3) Gaussian Regression under Fairness Constraints
M. D., Lucas De Lara, Élie Odin, François Bachoc
- (4) Characterization of the dual Cone of Monotone Functions
M. D., Fabrice Gamboa, Olivier Roustant

Gaussian Process Regression under Functional Constraints

Constrained GPs & prediction

Goal. Given observations $\{x_i, y_i\}_{i=1}^n := (\mathbf{x}, \mathbf{y})$, a functional constraint \mathcal{F} (e.g. monotonicity), and a GP $\{W(x) : x \in \mathcal{X}\} \sim \mathcal{GP}(0, k)$, construct:

$$W_{\mathcal{F}} := (W \mid \underbrace{W(\mathbf{x}) + \boldsymbol{\epsilon} = \mathbf{y}}_{\text{observations}}, \underbrace{W \in \mathcal{F}}_{\text{functional constraints}})$$

Constrained GPs & prediction

Goal. Given observations $\{x_i, y_i\}_{i=1}^n := (\mathbf{x}, \mathbf{y})$, a functional constraint \mathcal{F} (e.g. monotonicity), and a GP $\{W(x) : x \in \mathcal{X}\} \sim \mathcal{GP}(0, k)$, construct:

$$W_{\mathcal{F}} := (W \mid W(\mathbf{x}) + \epsilon = \mathbf{y}, W \in \mathcal{F})$$

Key idea. Project W onto some finite dim-space $\text{Span} \{\phi_i\}_{i=1}^m$:

$$\widetilde{W} = \Pi_{\Phi} W = \sum_{i=1}^m \xi_i \phi_i, \quad \boldsymbol{\xi} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}).$$

Constrained GPs & prediction

Goal. Given observations $\{x_i, y_i\}_{i=1}^n := (\mathbf{x}, \mathbf{y})$, a functional constraint \mathcal{F} (e.g. monotonicity), and a GP $\{W(x) : x \in \mathcal{X}\} \sim \mathcal{GP}(0, k)$, construct:

$$W_{\mathcal{F}} := (W \mid W(\mathbf{x}) + \epsilon = \mathbf{y}, W \in \mathcal{F})$$

Key idea. Project W onto some finite dim-space $\text{Span} \{\phi_i\}_{i=1}^m$:

$$\widetilde{W} = \Pi_{\Phi} W = \sum_{i=1}^m \xi_i \phi_i, \quad \boldsymbol{\xi} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}).$$

- $\boldsymbol{\xi}$ captures the randomness of \widetilde{W}

Constrained GPs & prediction

Goal. Given observations $\{x_i, y_i\}_{i=1}^n := (\mathbf{x}, \mathbf{y})$, a functional constraint \mathcal{F} (e.g. monotonicity), and a GP $\{W(x) : x \in \mathcal{X}\} \sim \mathcal{GP}(0, k)$, construct:

$$W_{\mathcal{F}} := (W \mid W(\mathbf{x}) + \epsilon = \mathbf{y}, W \in \mathcal{F})$$

Key idea. Project W onto some finite dim-space $\text{Span} \{\phi_i\}_{i=1}^m$:

$$\widetilde{W} = \Pi_{\Phi} W = \sum_{i=1}^m \xi_i \phi_i, \quad \boldsymbol{\xi} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}).$$

- $\boldsymbol{\xi}$ captures the randomness of \widetilde{W}

Prediction. Compute the law of the random vector:

$$\boldsymbol{\xi}_{\mathcal{F}} := (\boldsymbol{\xi} \mid \widetilde{W}(\mathbf{x}) + \epsilon = \mathbf{y}, \widetilde{W} \in \mathcal{F}),$$

and construct a predictor $\widehat{\boldsymbol{\xi}}_{\mathcal{F}} := \text{mod } \boldsymbol{\xi}_{\mathcal{F}}$.

Constrained GPs & prediction

Goal. Given observations $\{x_i, y_i\}_{i=1}^n := (\mathbf{x}, \mathbf{y})$, a functional constraint \mathcal{F} (e.g. monotonicity), and a GP $\{W(x) : x \in \mathcal{X}\} \sim \mathcal{GP}(0, k)$, construct:

$$W_{\mathcal{F}} := (W \mid W(\mathbf{x}) + \epsilon = \mathbf{y}, W \in \mathcal{F})$$

Key idea. Project W onto some finite dim-space $\text{Span}\{\phi_i\}_{i=1}^m$:

$$\widetilde{W} = \Pi_{\Phi} W = \sum_{i=1}^m \xi_i \phi_i, \quad \boldsymbol{\xi} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}).$$

- $\boldsymbol{\xi}$ captures the randomness of \widetilde{W}

Prediction. Compute the law of the random vector:

$$\boldsymbol{\xi}_{\mathcal{F}} := (\boldsymbol{\xi} \mid \widetilde{W}(\mathbf{x}) + \epsilon = \mathbf{y}, \widetilde{W} \in \mathcal{F}),$$

and construct a predictor $\widehat{\boldsymbol{\xi}}_{\mathcal{F}} := \text{mod } \boldsymbol{\xi}_{\mathcal{F}}$.  computation in $\mathcal{O}(m^3)$.

Block-Additive Gaussian Processes under Monotonicity Constraints - Statistics and Computing

Joint work with: Andrés F. López-Lopera, François Bachoc, Olivier Roustant, and Jérémy Rohmer

Contributions:

- Extending the framework of constrained GPs to the case of block-additive monotone functions:

$$y(x_1, \dots, x_D) = y_1(\mathbf{x}_{B_1}) + \dots + y_B(\mathbf{x}_{B_B}).$$

Block-Additive Gaussian Processes under Monotonicity Constraints - Statistics and Computing

Joint work with: Andrés F. López-Lopera, François Bachoc, Olivier Roustant, and Jérémy Rohmer

Contributions:

- Extending the framework of constrained GPs to the case of block-additive monotone functions:

$$y(x_1, \dots, x_D) = y_1(\mathbf{x}_{B_1}) + \dots + y_B(\mathbf{x}_{B_B}).$$

- A new algorithm *MaxMod*, that, given observations $\{\mathbf{x}_i, y(\mathbf{x}_i)\}_{i=1}^n$,
 - constructs a partition $\mathcal{P} := (B'_1, \dots, B'_{B'})$, together with
 - a monotone block-additive predictor \hat{y} of y .

Block-Additive Gaussian Processes under Monotonicity Constraints - Statistics and Computing

Joint work with: Andrés F. López-Lopera, François Bachoc, Olivier Roustant, and Jérémy Rohmer

Contributions:

- Extending the framework of constrained GPs to the case of block-additive monotone functions:

$$y(x_1, \dots, x_D) = y_1(\mathbf{x}_{B_1}) + \dots + y_B(\mathbf{x}_{B_B}).$$

- A new algorithm *MaxMod*, that, given observations $\{\mathbf{x}_i, y(\mathbf{x}_i)\}_{i=1}^n$,
 - constructs a partition $\mathcal{P} := (B'_1, \dots, B'_{B'})$, together with
 - a monotone block-additive predictor \hat{y} of y .
- Implementation of MaxMod for baGPs on the lineqGPR package:
<https://github.com/anfelopera/lineqGPR>.

Introduction to Fairness

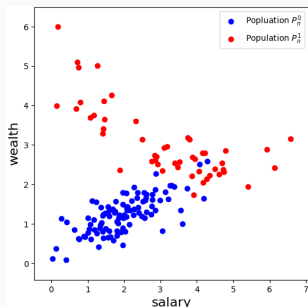
Example

Let $P_n = \{z_1, \dots, z_n\}$ be a population of n individuals, composed of the blue group P_n^0 and the red group P_n^1 s.t. $P_n = P_n^0 \cup P_n^1$.

Example

Let $P_n = \{z_1, \dots, z_n\}$ be a population of n individuals, composed of the blue group P_n^0 and the red group P_n^1 s.t. $P_n = P_n^0 \cup P_n^1$.

Each individual $z_i = (\mathbf{x}_i, s_i)$ is described by its **(wealth, salary)** = $\mathbf{x} \in \mathbb{R}^2$, and **color** $s \in \{0, 1\}$

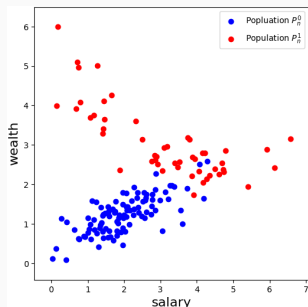


Plot of P_n

Example

Let $P_n = \{z_1, \dots, z_n\}$ be a population of n individuals, composed of the blue group P_n^0 and the red group P_n^1 s.t. $P_n = P_n^0 \cup P_n^1$.

Each individual $z_i = (\mathbf{x}_i, s_i)$ is described by its **(wealth, salary)** = $\mathbf{x} \in \mathbb{R}^2$, and **color** $s \in \{0, 1\}$ applies for a housing loan.

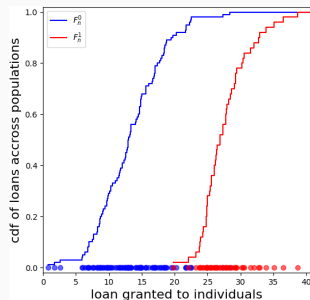


Plot of P_n

each z_i is attributed
a loan y_i

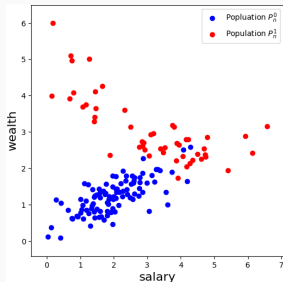
$$y_i = f^*(z_i) + \xi_i$$

ξ_i is some noise



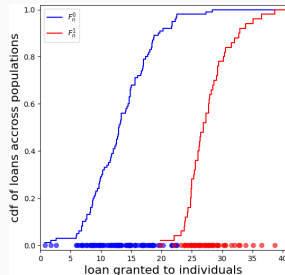
Plot of $\{y_i\}_{i=1}^n$ with CDFs

Ideas with examples for fairness



attributing of a loan
from f

$$z_i \mapsto f(z_i)$$



Constructing $f : (\text{wealth}, \text{salary}, \text{color}) \mapsto \text{loan}$

Ideas with examples for fairness

attributing of a loan
from f

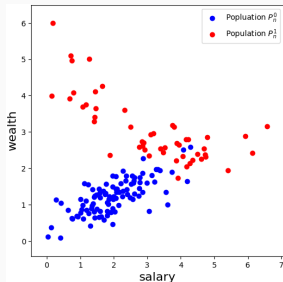
$$f(\cdot, \cdot, 0) = f(T, 1)$$

$$T \text{ s.t. } T(P_n^0) \simeq P_n^1$$

Constructing $f : (\text{wealth, salary, color}) \mapsto \text{loan}$

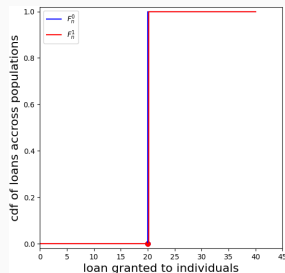
- **Statistical Parity.** The distribution of the loan $F_n^0 \simeq F_n^1$

Ideas with examples for fairness



attributing of a loan
from f

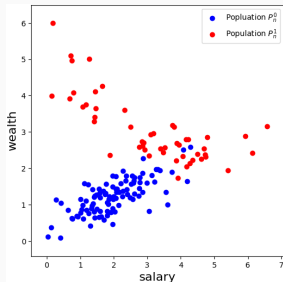
$$f : z_i \mapsto 20$$



Constructing $f : (\text{wealth}, \text{salary}, \text{color}) \mapsto \text{loan}$

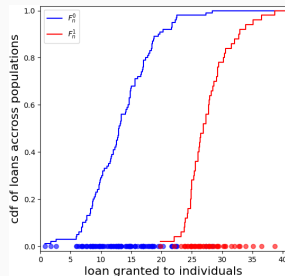
- **Statistical Parity.** The distribution of the loan $F_n^0 \simeq F_n^1$
- **No disparate treatment.** $f(\cdot, \cdot, 0) \simeq f(\cdot, \cdot, 1)$

Ideas with examples for fairness



attributing of a loan
from f

$$z_i \mapsto f(z_i)$$



Constructing $f : (\text{wealth}, \text{salary}, \text{color}) \mapsto \text{loan}$

- **Statistical Parity.** The distribution of the loan $F_n^0 \simeq F_n^1$
- **No disparate treatment.** $f(\cdot, \cdot, 0) \simeq f(\cdot, \cdot, 1)$
- **Accuracy.** $f(z_i) \simeq y_i, \forall i \in \{1, \dots, n\}$

Formalization of Fairness Notions

Let $(\Omega, \mathcal{A}, \mathbb{P})$ be a probability space, \mathcal{X} a Borel space, and $\mathcal{S} = \{0, 1\}$ a binary set (protected attribute), we call $\mathcal{Z} = \mathcal{X} \times \mathcal{S}$.

Let $Z := (X, S) : \Omega \rightarrow \mathcal{Z}$ be a random variable of law P and $Y : \Omega \rightarrow \mathbb{R}$ a real random variable.

We denote $P^s = \mathcal{L}((X, S) \mid S = s)$, $s \in \mathcal{S} = \{0, 1\}$.

Formalization of Fairness Notions

Let $(\Omega, \mathcal{A}, \mathbb{P})$ be a probability space, \mathcal{X} a Borel space, and $\mathcal{S} = \{0, 1\}$ a binary set (protected attribute), we call $\mathcal{Z} = \mathcal{X} \times \mathcal{S}$.

Let $Z := (X, S) : \Omega \rightarrow \mathcal{Z}$ be a random variable of law P and $Y : \Omega \rightarrow \mathbb{R}$ a real random variable.

We denote $P^s = \mathcal{L}((X, S) \mid S = s)$, $s \in \mathcal{S} = \{0, 1\}$.

Formalization of fairness

Let $f : \mathcal{Z} \rightarrow \mathbb{R}$, we say that f satisfies:

Statistical parity whenever $f(X, S) \perp\!\!\!\perp S \iff f_{\#}P^0 = f_{\#}P^1$

No disparate treatment whenever $f(\cdot, 0) = f(\cdot, 1)$

Formalization of Fairness Notions

Let $(\Omega, \mathcal{A}, \mathbb{P})$ be a probability space, \mathcal{X} a Borel space, and $\mathcal{S} = \{0, 1\}$ a binary set (protected attribute), we call $\mathcal{Z} = \mathcal{X} \times \mathcal{S}$.

Let $Z := (X, S) : \Omega \rightarrow \mathcal{Z}$ be a random variable of law P and $Y : \Omega \rightarrow \mathbb{R}$ a real random variable.

We denote $P^s = \mathcal{L}((X, S) \mid S = s)$, $s \in \mathcal{S} = \{0, 1\}$.

Formalization of fairness

Let $f : \mathcal{Z} \rightarrow \mathbb{R}$, we say that f satisfies:

Statistical parity whenever $f(X, S) \perp\!\!\!\perp S \iff f_{\#}P^0 = f_{\#}P^1$

No disparate treatment whenever $f(\cdot, 0) = f(\cdot, 1)$

Accuracy through quadratic risk.

We are looking at functions minimizing the *quadratic risk*

$$\mathbb{E} \left[|f(X, S) - Y|^2 \right], \quad \text{when } f \text{ is fair.}$$

Statistical Parity (S.P.): [Chzhen et al., 2020, Le Gouic et al., 2020]

$$\min \mathbb{E} \left[|f(X, S) - Y|^2 \right], \quad f \text{ satisfies S.P.}$$

S.P. & No Disparate Treatment (N.D.T.): [Divol and Gaucher, 2024]

$$\min \mathbb{E} \left[|f(X, S) - Y|^2 \right], \quad f \text{ satisfies S.P. and N.D.T.}$$

Fair GP for S.P. in Expectation: [Fitzsimons et al., 2019]

$$\text{Linear constraint: } \mathbb{E}[f(X, S = 0)] = \mathbb{E}[f(X, S = 1)]$$

Kernel Fair Regression: [Pérez-Suay et al., 2017, Li et al., 2019]

$$\min_{f \in \mathcal{H}} \tau \|f\|_{\mathcal{H}}^2 + \mathbb{E} [|f(X, S) - Y|^2] + \text{HSIC}(f(X, S), S)$$

Existence of a Convex C-loss for the Statistical Parity Constraint?

On the Nonconvexity of Push-Forward Constraints and Its Consequences in Machine Learning - SIMODS

Joint work with: Lucas De Lara, Alberto González-Sanz, and Virgil Foy

Contributions:

C is a subset of $\mathcal{G} := \{f : \mathbb{R}^d \rightarrow \mathbb{R}^p \text{ measurable}\}$

$J : \mathcal{G} \rightarrow [0, +\infty]$ is a C -loss if: $J^{-1}(\{0\}) = C$.

On the Nonconvexity of Push-Forward Constraints and Its Consequences in Machine Learning - SIMODS

Joint work with: Lucas De Lara, Alberto González-Sanz, and Virgil Foy

Contributions:

C is a subset of $\mathcal{G} := \{f : \mathbb{R}^d \rightarrow \mathbb{R}^p \text{ measurable}\}$

$J : \mathcal{G} \rightarrow [0, +\infty]$ is a C -loss if: $J^{-1}(\{0\}) = C$.

- If $J : \mathcal{G} \rightarrow [0, +\infty]$ is a convex function, then $f^{-1}(\{0\})$ is convex.
Hence, there exists no convex C -loss J if C is not convex.

On the Nonconvexity of Push-Forward Constraints and Its Consequences in Machine Learning - SIMODS

Joint work with: Lucas De Lara, Alberto González-Sanz, and Virgil Foy

Contributions:

C is a subset of $\mathcal{G} := \{f : \mathbb{R}^d \rightarrow \mathbb{R}^p \text{ measurable}\}$

$J : \mathcal{G} \rightarrow [0, +\infty]$ is a C -loss if: $J^{-1}(\{0\}) = C$.

- If $J : \mathcal{G} \rightarrow [0, +\infty]$ is a convex function, then $f^{-1}(\{0\})$ is convex. Hence, there exists no convex C -loss J if C is not convex.
- In general, the set of *equalizing maps* between P and $Q \in \mathcal{P}(\mathbb{R}^d)$, $\mathcal{E}(P, Q) := \{f : \mathbb{R}^d \rightarrow \mathbb{R}^p : f_{\#}P = f_{\#}Q\}$ is not convex.

On the Nonconvexity of Push-Forward Constraints and Its Consequences in Machine Learning - SIMODS

Joint work with: Lucas De Lara, Alberto González-Sanz, and Virgil Foy

Contributions:

C is a subset of $\mathcal{G} := \{f : \mathbb{R}^d \rightarrow \mathbb{R}^p \text{ measurable}\}$

$J : \mathcal{G} \rightarrow [0, +\infty]$ is a C -loss if: $J^{-1}(\{0\}) = C$.

- If $J : \mathcal{G} \rightarrow [0, +\infty]$ is a convex function, then $f^{-1}(\{0\})$ is convex. Hence, there exists no convex C -loss J if C is not convex.
- In general, the set of *equalizing maps* between P and $Q \in \mathcal{P}(\mathbb{R}^d)$, $\mathcal{E}(P, Q) := \{f : \mathbb{R}^d \rightarrow \mathbb{R}^p : f_{\#}P = f_{\#}Q\}$ is not convex.

Conclusion: In general, there exists no convex C -loss for the sets of *equalizing maps*.

Gaussian Process Regression under Fairness Constraints

Joint work with: François Bachoc, Lucas De Lara, and Élie Odin

Fair Gaussian Process

Empirical information. Let $(\mathbf{x}, \mathbf{s}, \mathbf{y}) = \{x_i, s_i, y_i\}_{i=1}^n$ be n i.i.d. realizations of (X, S, Y) . We write $Z = (X, S)$ and $z_i = (x_i, s_i)$. We split $\mathbf{z} = \{(x_i, s_i)\}_{i=1}^n$ in two: $\mathbf{z} = \mathbf{z}^0 \cup \mathbf{z}^1$, $\mathbf{z}^s = \{(x_i, s_i)\}_{i:s_i=s}$, $s \in \{0, 1\}$. Call $\hat{\mathbb{P}}_{\mathbf{z}^0}, \hat{\mathbb{P}}_{\mathbf{z}^1}$ the empirical measures associated to \mathbf{z}^0 and \mathbf{z}^1 .

Fair Gaussian Process

Empirical information. Let $(\mathbf{x}, \mathbf{s}, \mathbf{y}) = \{x_i, s_i, y_i\}_{i=1}^n$ be n i.i.d. realizations of (X, S, Y) . We write $Z = (X, S)$ and $z_i = (x_i, s_i)$. We split $\mathbf{z} = \{(x_i, s_i)\}_{i=1}^n$ in two: $\mathbf{z} = \mathbf{z}^0 \cup \mathbf{z}^1$, $\mathbf{z}^s = \{(x_i, s_i)\}_{i:s_i=s}$, $s \in \{0, 1\}$. Call $\hat{\mathbb{P}}_{\mathbf{z}^0}, \hat{\mathbb{P}}_{\mathbf{z}^1}$ the empirical measures associated to \mathbf{z}^0 and \mathbf{z}^1 .

δ -Fair Gaussian Process.

Let $\{W(z) : z \in \mathcal{Z}\} \sim \mathcal{GP}(0, k)$ and \mathcal{D} a distance on $\mathcal{P}(\mathbb{R})$.

We define the δ -fair process:

$$\widehat{W}_\delta := \left(W \mid \underbrace{W(\mathbf{z}) = \mathbf{y} + \epsilon}_{\text{Accuracy}}, \underbrace{\mathcal{D}\left(W_{\# \hat{\mathbb{P}}_{\mathbf{z}^0}}, W_{\# \hat{\mathbb{P}}_{\mathbf{z}^1}}\right) \leq \delta}_{\text{Statistical Parity}} \right),$$

where $\epsilon \sim \mathcal{N}(0, \tau \mathbf{I}_n)$ is a Gaussian noise independent of W .

Fair Gaussian Process

Empirical information. Let $(\mathbf{x}, \mathbf{s}, \mathbf{y}) = \{x_i, s_i, y_i\}_{i=1}^n$ be n i.i.d. realizations of (X, S, Y) . We write $Z = (X, S)$ and $z_i = (x_i, s_i)$. We split $\mathbf{z} = \{(x_i, s_i)\}_{i=1}^n$ in two: $\mathbf{z} = \mathbf{z}^0 \cup \mathbf{z}^1$, $\mathbf{z}^s = \{(x_i, s_i)\}_{i:s_i=s}$, $s \in \{0, 1\}$. Call $\hat{\mathbb{P}}_{\mathbf{z}^0}, \hat{\mathbb{P}}_{\mathbf{z}^1}$ the empirical measures associated to \mathbf{z}^0 and \mathbf{z}^1 .

δ -Fair Gaussian Process.

Let $\{W(z) : z \in \mathcal{Z}\} \sim \mathcal{GP}(0, k)$ and \mathcal{D} a distance on $\mathcal{P}(\mathbb{R})$.

We define the δ -fair process:

$$\widehat{W}_\delta := \left(W \mid \underbrace{W(\mathbf{z}) = \mathbf{y} + \epsilon}_{\text{Accuracy}}, \underbrace{\mathcal{D}\left(W_{\#}\hat{\mathbb{P}}_{\mathbf{z}^0}, W_{\#}\hat{\mathbb{P}}_{\mathbf{z}^1}\right) \leq \delta}_{\text{Statistical Parity}} \right),$$

where $\epsilon \sim \mathcal{N}(0, \tau \mathbf{I}_n)$ is a Gaussian noise independent of W .

Remark. $\mathcal{D}\left(W_{\#}\hat{\mathbb{P}}_{\mathbf{z}^0}, W_{\#}\hat{\mathbb{P}}_{\mathbf{z}^1}\right)$ is a function of $W(\mathbf{z})$.

Construction of Fair Predictions

Definition. Maximum a Posteriori (MAP) prediction

The MAP prediction, $\mathcal{Z}^p \rightarrow \mathbb{R}^p$ is defined for every element $\mathbf{t} \in \mathcal{Z}^p$ as

$$\text{pred}(\mathbf{t}) := \pi_{[p]}(\text{mod } \widehat{W}_\delta(\mathbf{t}, \mathbf{z})).$$

Construction of Fair Predictions

Definition. Maximum a Posteriori (MAP) prediction

The MAP prediction, $\mathcal{Z}^p \rightarrow \mathbb{R}^p$ is defined for every element $\mathbf{t} \in \mathcal{Z}^p$ as

$$\text{pred}(\mathbf{t}) := \pi_{[p]}(\text{mod } \widehat{W}_\delta(\mathbf{t}, \mathbf{z})).$$

Theorem. Consistency of the MAP

There exists $\widehat{f}_\delta : \mathcal{Z} \rightarrow \mathbb{R}$ s.t. $\forall \mathbf{t} \in \mathcal{Z}^p$:

$$\text{pred}(\mathbf{t}) = \widehat{f}_\delta(\mathbf{t}) = k(\mathbf{t}, \mathbf{z})k(\mathbf{z}, \mathbf{z})^{-1}\widehat{\mathbf{w}}_\delta, \quad (1)$$

where k is the kernel of W

Construction of Fair Predictions

Definition. Maximum a Posteriori (MAP) prediction

The MAP prediction, $\mathcal{Z}^p \rightarrow \mathbb{R}^p$ is defined for every element $\mathbf{t} \in \mathcal{Z}^p$ as

$$\text{pred}(\mathbf{t}) := \pi_{[p]}(\text{mod } \widehat{W}_\delta(\mathbf{t}, \mathbf{z})).$$

Theorem. Consistency of the MAP

There exists $\widehat{f}_\delta : \mathcal{Z} \rightarrow \mathbb{R}$ s.t. $\forall \mathbf{t} \in \mathcal{Z}^p$:

$$\text{pred}(\mathbf{t}) = \widehat{f}_\delta(\mathbf{t}) = k(\mathbf{t}, \mathbf{z})k(\mathbf{z}, \mathbf{z})^{-1}\widehat{\mathbf{w}}_\delta, \quad (1)$$

where k is the kernel of W and $\widehat{\mathbf{w}}_\delta$ is a solution of

$$\begin{aligned} \min_{\mathbf{w}=(\mathbf{w}^0, \mathbf{w}^1) \in \mathbb{R}^n} \quad & \|\mathbf{w} - \boldsymbol{\mu}\|_{\Gamma^{-1}}^2 \\ \text{subject to} \quad & \mathcal{D}(\widehat{\mathbb{P}}_{\mathbf{w}^0}, \widehat{\mathbb{P}}_{\mathbf{w}^1}) \leq \delta, \end{aligned} \quad (\text{Pb}_\delta)$$

$(W(\mathbf{z}) \mid W(\mathbf{z}) = \mathbf{y} + \epsilon) \sim \mathcal{N}(\boldsymbol{\mu}, \Gamma)$ and $\widehat{\mathbb{P}}_{\mathbf{w}^s}$ is the empirical probability measure associated to \mathbf{w}^s for $s \in \{0, 1\}$.

Computation of the MAP predictor

The MAP problem. We aim to solve:

$$\begin{aligned} \min_{\mathbf{w}=(\mathbf{w}^0, \mathbf{w}^1) \in \mathbb{R}^n} \quad & \|\mathbf{w} - \boldsymbol{\mu}\|_{\Gamma^{-1}}^2 \\ \text{subject to} \quad & \mathcal{D}(\hat{\mathbb{P}}_{\mathbf{w}^0}, \hat{\mathbb{P}}_{\mathbf{w}^1}) \leq \delta, \end{aligned} \tag{Pb}_\delta$$

Computation of the MAP predictor

The MAP problem. We aim to solve:

$$\begin{aligned} \min_{\mathbf{w}=(\mathbf{w}^0, \mathbf{w}^1) \in \mathbb{R}^n} \quad & \|\mathbf{w} - \boldsymbol{\mu}\|_{\Gamma^{-1}}^2 \\ \text{subject to} \quad & \mathcal{D}(\hat{\mathbb{P}}_{\mathbf{w}^0}, \hat{\mathbb{P}}_{\mathbf{w}^1}) \leq \delta, \end{aligned} \quad (\text{Pb}_\delta)$$

Penalized problem. In practice, we solve the penalized problem:

$$\min_{\mathbf{w}=(\mathbf{w}^0, \mathbf{w}^1) \in \mathbb{R}^n} \quad \|\mathbf{w} - \boldsymbol{\mu}\|_{\Gamma^{-1}}^2 + \lambda \mathcal{D}(\hat{\mathbb{P}}_{\mathbf{w}^0}, \hat{\mathbb{P}}_{\mathbf{w}^1}). \quad (\text{Pb}_\lambda)$$

Computation of the MAP predictor

The MAP problem. We aim to solve:

$$\begin{aligned} \min_{\mathbf{w}=(\mathbf{w}^0, \mathbf{w}^1) \in \mathbb{R}^n} \quad & \|\mathbf{w} - \boldsymbol{\mu}\|_{\Gamma^{-1}}^2 \\ \text{subject to} \quad & \mathcal{D}(\hat{\mathbb{P}}_{\mathbf{w}^0}, \hat{\mathbb{P}}_{\mathbf{w}^1}) \leq \delta, \end{aligned} \quad (\text{Pb}_\delta)$$

Penalized problem. In practice, we solve the penalized problem:

$$\min_{\mathbf{w}=(\mathbf{w}^0, \mathbf{w}^1) \in \mathbb{R}^n} \quad \|\mathbf{w} - \boldsymbol{\mu}\|_{\Gamma^{-1}}^2 + \lambda \mathcal{D}(\hat{\mathbb{P}}_{\mathbf{w}^0}, \hat{\mathbb{P}}_{\mathbf{w}^1}). \quad (\text{Pb}_\lambda)$$

Correspondence between (Pb_λ) and (Pb_δ)

For all $\lambda > 0$:

- $\hat{\mathbf{w}}_\lambda = (\hat{\mathbf{w}}_\lambda^0, \hat{\mathbf{w}}_\lambda^1) \text{ sol } (\text{Pb}_\lambda) \implies \hat{\mathbf{w}}_\lambda \text{ sol } (\text{Pb}_\delta) \text{ if } \delta = \mathcal{D}(\hat{\mathbb{P}}_{\hat{\mathbf{w}}_\lambda^0}, \hat{\mathbb{P}}_{\hat{\mathbf{w}}_\lambda^1})$

Computation of the MAP predictor

The MAP problem. We aim to solve:

$$\begin{aligned} \min_{\mathbf{w}=(\mathbf{w}^0, \mathbf{w}^1) \in \mathbb{R}^n} \quad & \|\mathbf{w} - \boldsymbol{\mu}\|_{\Gamma^{-1}}^2 \\ \text{subject to} \quad & \mathcal{D}(\hat{\mathbb{P}}_{\mathbf{w}^0}, \hat{\mathbb{P}}_{\mathbf{w}^1}) \leq \delta, \end{aligned} \quad (\text{Pb}_\delta)$$

Penalized problem. In practice, we solve the penalized problem:

$$\min_{\mathbf{w}=(\mathbf{w}^0, \mathbf{w}^1) \in \mathbb{R}^n} \quad \|\mathbf{w} - \boldsymbol{\mu}\|_{\Gamma^{-1}}^2 + \lambda \mathcal{D}(\hat{\mathbb{P}}_{\mathbf{w}^0}, \hat{\mathbb{P}}_{\mathbf{w}^1}). \quad (\text{Pb}_\lambda)$$

Correspondence between (Pb_λ) and (Pb_δ)

For all $\lambda > 0$:

- $\hat{\mathbf{w}}_\lambda = (\hat{\mathbf{w}}_\lambda^0, \hat{\mathbf{w}}_\lambda^1) \text{ sol } (\text{Pb}_\lambda) \implies \hat{\mathbf{w}}_\lambda \text{ sol } (\text{Pb}_\delta) \text{ if } \delta = \mathcal{D}(\hat{\mathbb{P}}_{\hat{\mathbf{w}}_\lambda^0}, \hat{\mathbb{P}}_{\hat{\mathbf{w}}_\lambda^1})$
- $\mathcal{D}(\hat{\mathbb{P}}_{\hat{\mathbf{w}}_\lambda^0}, \hat{\mathbb{P}}_{\hat{\mathbf{w}}_\lambda^1}) \leq \|\boldsymbol{\mu}\|_{\Gamma^{-1}}^2 / \lambda$

Hence, the empirical statistical parity is controlled by λ .

A Kimeldorf-Wahba Correspondence

Let $\widehat{\mathbf{w}}_\lambda$ be a solution of (Pb_λ) , define

$$\widehat{f}_\lambda(\mathbf{t}) = k(\mathbf{t}, \mathbf{z})k(\mathbf{z}, \mathbf{z})^{-1}\widehat{\mathbf{w}}_\lambda.$$

A Kimeldorf-Wahba Correspondence

Let $\widehat{\mathbf{w}}_\lambda$ be a solution of (Pb_λ) , define

$$\widehat{f}_\lambda(\mathbf{t}) = k(\mathbf{t}, \mathbf{z})k(\mathbf{z}, \mathbf{z})^{-1}\widehat{\mathbf{w}}_\lambda.$$

Theorem

The function \widehat{f}_λ , is solution of the Tikhonov-regularized minimization problem penalized with fairness constraint:

$$\min_{f \in \mathcal{H}} \underbrace{\tau \|f\|_{\mathcal{H}}^2}_{\text{noise } \epsilon} + \underbrace{\sum_{i=1}^n |f(z_i) - y_i|^2}_{\text{empirical } L^2 \text{ risk}} + \underbrace{\lambda \mathcal{D}(f_{\#}\widehat{\mathbb{P}}_{\mathbf{z}^0}, f_{\#}\widehat{\mathbb{P}}_{\mathbf{z}^1})}_{\text{fairness penalty}}. \quad (\text{Pb}_{\mathcal{H}, \lambda})$$

where \mathcal{H} is the RKHS generated by the kernel k of the GP W .

A Kimeldorf-Wahba Correspondence

Let $\widehat{\mathbf{w}}_\lambda$ be a solution of (Pb_λ) , define

$$\widehat{f}_\lambda(\mathbf{t}) = k(\mathbf{t}, \mathbf{z})k(\mathbf{z}, \mathbf{z})^{-1}\widehat{\mathbf{w}}_\lambda.$$

Theorem

The function \widehat{f}_λ , is solution of the Tikhonov-regularized minimization problem penalized with fairness constraint:

$$\min_{f \in \mathcal{H}} \underbrace{\tau \|f\|_{\mathcal{H}}^2}_{\text{noise } \epsilon} + \underbrace{\sum_{i=1}^n |f(z_i) - y_i|^2}_{\text{empirical } L^2 \text{ risk}} + \underbrace{\lambda \mathcal{D}(f_{\#}\widehat{\mathbb{P}}_{\mathbf{z}^0}, f_{\#}\widehat{\mathbb{P}}_{\mathbf{z}^1})}_{\text{fairness penalty}}. \quad (\text{Pb}_{\mathcal{H}, \lambda})$$

where \mathcal{H} is the RKHS generated by the kernel k of the GP W .

Remark. This is the fair kernel regression problem in [Pérez-Suay et al., 2017].

Asymptotic properties of the regressions functions

Let $(z_i, y_i)_{i \geq 1}$ be a sequence of i.i.d. realizations of $(Z, Y) : \Omega \rightarrow \mathcal{Z} \times \mathbb{R}$.

For all n , let $\mathbf{z}_n = \{z_i\}_{i=1}^n = \mathbf{z}_n^0 \cup \mathbf{z}_n^1 \in \mathcal{Z}^n$, $\tau_n, \lambda_n > 0$

Let \mathcal{G}_n be the set of solutions of:

$$\min_{f \in \mathcal{H}} \tau_n \|f\|_{\mathcal{H}}^2 + \frac{1}{n} \sum_{i=1}^n |f(z_i) - y_i|^2 + \lambda_n \mathcal{D} \left(f_{\#} \widehat{\mathbb{P}}_{\mathbf{z}_n^0}, f_{\#} \widehat{\mathbb{P}}_{\mathbf{z}_n^1} \right) \quad (\text{Pb}_{\mathcal{H},n})$$

Asymptotic properties of the regressions functions

Let $(z_i, y_i)_{i \geq 1}$ be a sequence of i.i.d. realizations of $(Z, Y) : \Omega \rightarrow \mathcal{Z} \times \mathbb{R}$.

For all n , let $\mathbf{z}_n = \{z_i\}_{i=1}^n = \mathbf{z}_n^0 \cup \mathbf{z}_n^1 \in \mathcal{Z}^n$, $\tau_n, \lambda_n > 0$

Let \mathcal{G}_n be the set of solutions of:

$$\min_{f \in \mathcal{H}} \tau_n \|f\|_{\mathcal{H}}^2 + \frac{1}{n} \sum_{i=1}^n |f(z_i) - y_i|^2 + \lambda_n \mathcal{D} \left(f_{\#} \widehat{\mathbb{P}}_{\mathbf{z}_n^0}, f_{\#} \widehat{\mathbb{P}}_{\mathbf{z}_n^1} \right) \quad (\text{Pb}_{\mathcal{H},n})$$

Theorem. Convergence of the solution set

\mathcal{G}_n converges a.s. to \mathcal{G}^* in Hausdorff distance.

where \mathcal{G}^* is the set of solutions of the problem

$$\min_{f \in \mathcal{H}} \tau \|f\|_{\mathcal{H}}^2 + \mathbb{E} [|f(Z) - Y|^2] + \lambda \mathcal{D} (f_{\#} P^0, f_{\#} P^1). \quad (\text{Pb}_{\mathcal{H},*})$$

If $\tau_n \rightarrow \tau > 0$ and $\lambda_n \rightarrow \lambda$. Recall $P^s = \mathcal{L}(Z | S = s)$.

Parametrization for controlling correlation of GPs

Parametrization choices

Assume $\{W(x, s) : (x, s) \in \mathcal{X} \times \mathcal{S}\} \sim \mathcal{GP}(0, k^\rho)$ where

$$k^\rho((x, s), (x', s')) = k_{\mathcal{X}}(x, x') \cdot k_{\mathcal{S}}^\rho(s, s').$$

Where $k_{\mathcal{S}}^\rho$ is defined as:

$$\begin{bmatrix} k_{\mathcal{S}}^\rho(\mathbf{0}, \mathbf{0}) & k_{\mathcal{S}}^\rho(\mathbf{0}, \mathbf{1}) \\ k_{\mathcal{S}}^\rho(\mathbf{1}, \mathbf{0}) & k_{\mathcal{S}}^\rho(\mathbf{1}, \mathbf{1}) \end{bmatrix} = \begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix},$$

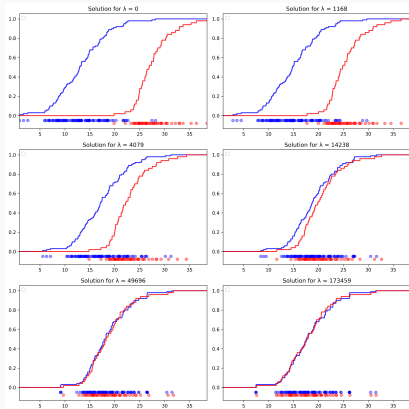
then, $\rho = \text{Cor}(W(x, \mathbf{0}), W(x, \mathbf{1})) \quad \forall x \in \mathcal{X}$.

Consequences of the choice of ρ .

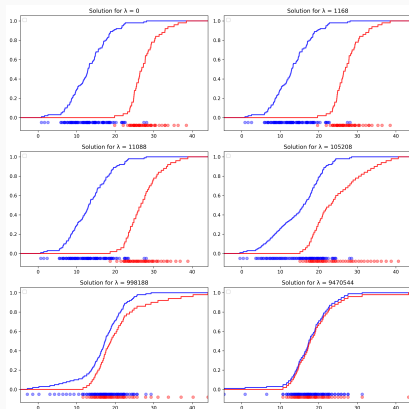
- When $\rho = 0$, $W(x, \mathbf{0}) \perp W(y, \mathbf{1})$, $\forall x, y \in \mathcal{X}$.
- When $\rho = 1$, $W(x, \mathbf{0}) = W(x, \mathbf{1})$, $\forall x \in \mathcal{X}$ a.s.
- Control on the N.D.T. : $\left\| \widehat{f}_\rho(\cdot, \mathbf{0}) - \widehat{f}_\rho(\cdot, \mathbf{1}) \right\|_{\mathcal{H}_x} \leq |1 - \rho|A$.

Numerical Results

Numerical results on the example for different value of λ



Example of fair predictor with Sinkhorn divergence



Example of fair predictors with MMD distance

Numerical results of CRIME dataset

Methodology. Split the dataset (70% training, 30% testing).
Computation of Mean Square Error (MSE) and Kolmogorov-Smirnov (KS): $\|\mathcal{F}^0 - \mathcal{F}^1\|_\infty$ on the test set and comparison of the results with the state of the art [Chzhen et al., 2020].

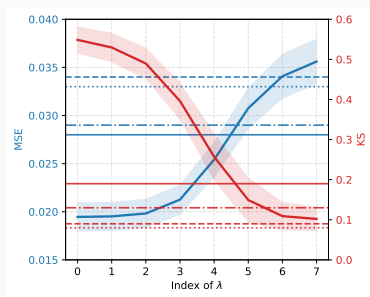


Figure 1:

$$\rho = 0 \iff W(\cdot, 0) \perp W(\cdot, 1)$$

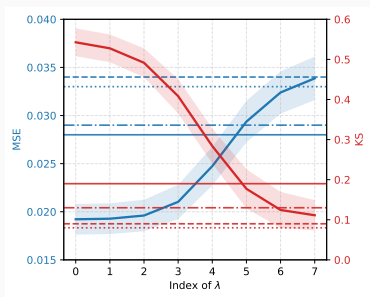


Figure 2:

$$\rho = 1 \iff W(\cdot, 0) = W(\cdot, 1)$$

Summary of Results & Conclusion

Summary of Results

MAP/RKHS correspondence

1. We developed a method to construct a “fair” predictor with a MAP method from a δ -fair GP.
2. The *fair GP regression* corresponds to *kernel fair regression*.
3. Asymptotically, the MAP predictors converge to the solutions of a regularized constrained problem.

Summary of Results

MAP/RKHS correspondence

1. We developed a method to construct a “fair” predictor with a MAP method from a δ -fair GP.
2. The *fair GP regression* corresponds to *kernel fair regression*.
3. Asymptotically, the MAP predictors converge to the solutions of a regularized constrained problem.

Control on the fairness constraint

- 4 We can control the **empirical statistical parity** constraint with parameter λ in (Pb_λ) .
- 5 We can control the **no disparate treatment** through kernel parametrization k_S^ρ .

Summary of Results

MAP/RKHS correspondence

1. We developed a method to construct a “fair” predictor with a MAP method from a δ -fair GP.
2. The *fair GP regression* corresponds to *kernel fair regression*.
3. Asymptotically, the MAP predictors converge to the solutions of a regularized constrained problem.






Control on the fairness constraint

- 4 We can control the **empirical statistical parity** constraint with parameter λ in (Pb_λ) .
- 5 We can control the **no disparate treatment** through kernel parametrization k_S^ρ .

Uncertainty Quantification

On our predictions? On the fairness?

Bibliography

-  Chzhen, E., Denis, C., Hebiri, M., Oneto, L., and Pontil, M. (2020).
Fair regression with Wasserstein barycenters.
Advances in Neural Information Processing Systems, 33:7321–7331.
-  Divol, V. and Gaucher, S. (2024).
Demographic parity in regression and classification within the unawareness framework.
arXiv preprint arXiv:2409.02471.
-  Fitzsimons, J., Ali, A. A., Osborne, M., and Roberts, S. (2019).
A General Framework for Fair Regression.
Entropy, 21(8):741.
-  Le Gouic, T., Loubes, J.-M., and Rigollet, P. (2020).
Projection to Fairness in Statistical Learning.
-  Li, Z., Perez-Suay, A., Camps-Valls, G., and Sejdinovic, D. (2019).
Kernel dependence regularizers and Gaussian processes with applications to algorithmic fairness.
arXiv preprint arXiv:1911.04322.

Thank you! Questions?
